## CISC 271 Class 21

## PCA - Matrix Algebra and Dimensionality Reduction

Text Correspondence: ~

## Main Concepts:

- Scatter matrix
- PCA for dimensionality reduction
- Clustering of reduced data

Sample Problem, Machine Inference: How can PCA reduce the number of variables in data?

We previously presented principal components analysis (PCA) by exploring the covariance of data. A data matrix $A$ has a mean row-like matrix $\bar{A}$ and a zero-mean matrix $M$ that we can re-state as

$$
\begin{align*}
\bar{A} & \stackrel{\text { def }}{=}  \tag{21.1}\\
M & \frac{1}{m} \overrightarrow{1}^{T} A \\
M & \stackrel{\text { def }}{=} A-\overrightarrow{1} \bar{A}
\end{align*}
$$

Using the sample variance, the matrix $M$ in Equation 21.1 can be used to define the covariance matrix as

$$
\begin{equation*}
B=\left(\frac{1}{m-1}\right) M^{T} M \tag{21.2}
\end{equation*}
$$

The eigenvalues of the matrix $B$ in Equation 21.2 are non-negative and can be interpreted as describing statistical relationships of columns of the data matrix $A$. A closely related matrix is the scatter matrix. This is a symmetric positive semidefinite matrix that we will define as

$$
\begin{equation*}
S \stackrel{\text { def }}{=} M^{T} M \tag{21.3}
\end{equation*}
$$

Because the matrix $S$ in Equation 21.3 is a scalar multiple of the matrix $B$ in Equation 21.2, the matrix $S$ and the matrix $B$ have the same eigenvectors. The eigenvalues are also scalar multiples, that is, the $j^{\text {th }}$ eigenvalue of $S$ is $m-1$ times the $j^{\text {th }}$ eigenvalue of $B$.

We can think of the scatter matrix $S$ as the weighted covariance. The scatter matrix can be used in place of the covariance matrix when the number of observations are of concern, or when the interpretation of the eigenvalues of the covariance matrix are unaffected by the factor of $m-1$.

Henceforth, we will work with PCA using the scatter matrix $B$ instead of using the covariance matrix $B$.

### 21.1 Scatter Matrix, SVD, and PCA

Using terminology that is common in PCA literature, we can say the PCA loading vectors $\vec{v}_{j}$ of the covariance matrix $B$ are the same as the eigenvectors of the scatter matrix $S$.

Consider the SVD of the zero-mean matrix, which is $M=U \Sigma V^{T}$. We can write the scatter matrix as

$$
\begin{align*}
S & =M^{T} M \\
& =\left[U \Sigma V^{T}\right]^{T} U \Sigma V^{T} \\
& =V \Sigma^{T} \Sigma V^{T} \\
& =V \Sigma^{2} V^{T} \tag{21.4}
\end{align*}
$$

As an aside, because the covariance matrix $B$ is a scalar multiple of the scatter matrix $S$, we can use Equation 21.4 to write the spectral decomposition for PCA as

$$
\begin{align*}
B & =\left(\frac{1}{m-1}\right) M^{T} M \\
& =V\left[\frac{\Sigma^{2}}{m-1}\right] V^{T} \tag{21.5}
\end{align*}
$$

Equation 21.4 provides us with the spectral decomposition of the scatter matrix $S$. We can observe that the eigenvectors $\vec{v}_{j}$ are the right singular vectors of the data matrix $M$. The scatter matrix is symmetric and it is positive definite if and only if the zero-mean matrix $M$ is full rank, that is, if and only if $M$ has each singular value greater than zero. Because of Equation 21.5, these observation also apply to the covariance matrix $B$.

For PCA, we defined the first score vector as the product of the zero-mean matrix $M$ and the first loading vector $\vec{v}_{1}$, that is, as $\vec{z}_{1}=M \vec{v}_{1}$. The right singular matrix $V$ of the zero-mean matrix $M$ has orthonormal columns, which can be summarized as

$$
\begin{align*}
\vec{v}_{j}^{T} \vec{v}_{j} & =1  \tag{21.6}\\
\vec{v}_{j}^{T} \vec{v}_{i \neq j} & =0
\end{align*}
$$

Using the properties of Equation 21.6, we have

$$
V^{T} \vec{v}_{1}=\left[\begin{array}{c}
1  \tag{21.7}\\
0 \\
\vdots \\
0
\end{array}\right] \quad \Rightarrow \quad \Sigma V^{T} \vec{v}_{1}=\left[\begin{array}{c}
\sigma_{1} \\
0 \\
\vdots \\
0
\end{array}\right]
$$

The first score vector $\vec{z}_{1}$, using the SVD of the zero-mean matrix $M$ and Equation 21.7, can be written as

$$
\begin{align*}
\vec{z}_{1} & =M \vec{v}_{1} \\
& =U \Sigma V^{T} \vec{v}_{1} \\
& =\sigma_{1} \vec{u}_{1} \tag{21.8}
\end{align*}
$$

Using the same reasoning of Equation 21.8 for each score, we can write the $j^{\text {th }}$ score vector of the data matrix $A$ as

$$
\begin{equation*}
\vec{z}_{j}=\sigma_{j} \vec{u}_{j} \tag{21.9}
\end{equation*}
$$

### 21.2 PCA and Low-Rank Approximation

Recall, from a previous class, that the optimal approximation of a matrix is a truncation of the Eckart-Young series. This series, for the zero-mean matrix $M$ that has rank $r$, is

$$
\begin{align*}
M & =U \Sigma V^{T} \\
& =\left[\sigma_{1} \vec{u}_{1}\right] \vec{v}_{1}^{T}+\left[\sigma_{2} \vec{u}_{2}\right] \vec{v}_{2}^{T}+\cdots+\left[\sigma_{r} \vec{u}_{r}\right] \vec{v}_{r}^{T} \tag{21.10}
\end{align*}
$$

Consider the first $p \leq r$ scores of the data in the original matrix $A$. These are the vectors $\vec{z}_{j}$ of Equation 21.9; we can gather these into a matrix $Z_{p}$. The first $p$ eigenvectors of the scatter matrix $S$ are the first $p$ right singular vectors of the zero-mean matrix $M$, which we can gather into a matrix $V_{p}$. The $p$ PCA scores and the $p$ PCA loading vectors - which are the right singular vectors of $M-$ are related by Equation 21.10 as

$$
\begin{align*}
Z_{p} & =M V_{p} \\
& =\left[\begin{array}{llll}
\sigma_{1} \vec{u}_{1} & \sigma_{2} \vec{u}_{2} & \cdots & \sigma_{p} \vec{u}_{p}
\end{array}\right] \tag{21.11}
\end{align*}
$$

An orthonormal basis for the column space of $Z_{p}$ is, simply, to divide each column by the nonzero singular value $\sigma_{j}$. Each basis vector is $\vec{u}_{j}$, which is a left singular vector of the zero-mean matrix $M$; together, an orthonormal basis for the first $p$ PCA scores are the first $p$ columns of the left singular matrix $U$, which we can abbreviate as $U_{p}$. This leads us to a remarkable observation:

The first p PCA scores are an optimal approximation of a $\mathrm{p}-\mathrm{D}$ vector space for the zero-mean data

### 21.3 PCA for Dimensionality Reduction

PCA is commonly used to reduce the dimensionality of data. Dimensionality reduction is the process of transforming a linear problem that has $n$ variables to a smaller problem that has $p<n$ variables. We have derived three equivalent methods to perform dimensionality reduction:

- Compute the first $p$ scores of the data matrix $A$ by using PCA, which is $Z_{p}$
- Compute the first $p$ eigenvectors of the scatter matrix $S$ as $V_{p}$ and compute $U_{p} \Sigma_{p}=M V_{p}$
- Compute the SVD of the zero-mean matrix $M$ and use the first $p$ left singular vectors $U_{p}$

These equivalent methods of computing dimensionality reduction have the same effect. For the original $n$-D vector space of the zero-mean data, which we can abbreviate as $\mathbb{U}_{n}$, the methods find an orthogonal basis for a $p$ - D vector space, which we can abbreviate as $\mathbb{U}_{p}$, with the important property that

$$
p<n \text { and } \mathbb{U}_{p} \subset \mathbb{U}_{n}
$$

The latter property holds because the orthogonal left singular matrix $U$ is an orthonormal basis for the column space of the zero-mean matrix $M$, so the first $n$ columns $U_{n}$ are a basis for $\mathbb{U}_{n}$.

Clustering algorithms have been found, from much empirical experience, to often perform poorly on high-dimensional data. A commonly used solution is to transform the data to a lowerdimensional vector space, and then to perform clustering by method such as $k$-means.

